Review of Probability Theory

Zahra Koochak and Jeremy Irvin

Sample Space Ω

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 $\{HH, HT, TH, TT\}$

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Event $A \subseteq \Omega$

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 $\{HH,HT\},$

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Event Space ${\mathcal F}$

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Probability Measure $P : \mathcal{F} \to \mathbb{R}$

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Probability Measure $P : \mathcal{F} \to \mathbb{R}$ $P(A) \ge 0 \quad \forall A \in \mathcal{F}$

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Probability Measure $P : \mathcal{F} \to \mathbb{R}$ $P(A) \ge 0 \quad \forall A \in \mathcal{F}$ $P(\Omega) = 1$

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Event $A \subseteq \Omega$

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Event Space \mathcal{F}

Probability Measure $P : \mathcal{F} \to \mathbb{R}$ $P(A) \ge 0 \quad \forall A \in \mathcal{F}$ $P(\Omega) = 1$

If $A_1, A_2, ...$ disjoint set of events $(A_i \cap A_j = \emptyset$ when $i \neq j)$, then

$$P\left(\bigcup_i A_i\right) = \sum_i P(A_i)$$

Let *B* be any event such that $P(B) \neq 0$.

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 $A \perp B$ if and only if $P(A \cap B) = P(A)P(B)$

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 if and only if $P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(A)P(B)}{P(B)} = P(A)$

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A **RV** is $X : \Omega \to \mathbb{R}$

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 $Val(X) := X(\Omega)$

 $\omega_0 = HHHTHTTHTT$

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of heads: $X(\omega_0) = 5$

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 $Val(X) := X(\Omega)$

 $Val(X) = \{0, 1, ..., 10\}$

$$F_X:\mathbb{R}\to[0,1]$$

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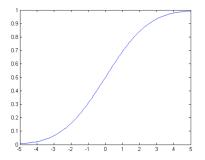
$$F_X(x) = P(X \leq x)$$

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 $\frac{\text{Probability Density Function (PDF)}}{f_X: \mathbb{R} \to \mathbb{R}}$

 $f_X(x) := \frac{d}{dx} F_X(x)$

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 $\frac{\text{Probability Density Function (PDF)}}{f_X : \mathbb{R} \to \mathbb{R}}$ $f_X(x) := \frac{d}{dx} F_X(x)$ $f_X(x) \neq P(X = x)$

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Let X be a discrete RV with PMF p_X .

$$\mathbb{E}[g(X)] := \sum_{x \in Val(X)} g(x) p_X(x)$$

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Variance

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Variance

$$Var(X) := \mathbb{E}[(X - \mathbb{E}[X])^2]$$

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Variance

$$Var(X) := \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

Example Distributions

Distribution	PDF or PMF	Mean	Variance
Bernoulli(p)	$\begin{cases} p, & \text{if } x = 1\\ 1-p, & \text{if } x = 0. \end{cases}$	p	p(1-p)
Binomial(n, p)	$\binom{n}{k} p^k (1-p)^{n-k}$ for $k = 0, 1,, n$	np	np(1-p)
Geometric(p)	$p(1-p)^{k-1}$ for $k = 1, 2,$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Poisson(λ)	$\frac{e^{-\lambda}\lambda^k}{k!}$ for $k = 0, 1,$	λ	λ
Uniform(a, b)	$rac{1}{b-a}$ for all $x\in(a,b)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Gaussian (μ, σ^2)	$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}} \text{ for all } x \in (-\infty,\infty)$	μ	σ^2
Exponential(λ)	$\lambda e^{-\lambda x}$ for all $x \ge 0, \lambda \ge 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$

Two Random Variables

Bivariate CDF

$$F_{XY}(x,y) = P(X \le x, Y \le y)$$

Bivariate PMF

$$p_{XY}(x,y) = P(X = x, Y = y)$$

Marginal PMF

$$p_X(x) = \sum_y p_{XY}(x, y)$$

Bivariate PDF

$$f_{XY}(x,y) = \frac{\partial^2 F_{XY}(x,y)}{\partial x \partial y}$$

Marginal PDF

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) dy$$

Bayes' Theorem

- Given the conditional probability of an event P(x|y)
- Want to find the "reverse" conditional probability, P(y|x)

$$P(y|x) = \frac{P(x|y)P(y)}{P(x)}$$

where: $P(x) = \sum_{y' \in value \ y} P(x|y')P(y')$

X and Y are continuous

$$f(y|x) = \frac{f(x|y)f(y)}{f(x)}$$

where: $f(x) = \int_{y' \in value \ y} f(x|y')f(y')dy'$

Example for Bayes Rule

You randomly choose a treasure chest to open, and then randomly choose a coin from that treasure chest. If the coin you choose is gold, then what is the probability that you choose chest A?

Independence

Two random variables X and Y are independent if:

$$p_{XY}(x,y) = p_X(x)p_Y(y)$$

$$\blacktriangleright p_{Y|X}(x,y) = P_Y(y)$$

For continuous random variables:

$$p_{XY}(x,y) \rightarrow f_{XY}(x,y)$$

Example for independent random variables

Spin a spinner numbered 1 to 7, and toss a coin. What is the probability of getting an odd. number on the spinner and a tail on the coin?

$$p_{XY}(x,y) = p_X(x)p_Y(y) = \frac{1}{2} \times \frac{4}{7} = \frac{2}{7}$$



Expectation

- X, Y :Two continuous random variables
- \blacktriangleright g , R2 \rightarrow R : A function of X and Y

$$E(g(x,y)) = \int_{x \in Val(x)} \int_{y \in Val(y)} g(x,y) f_{XY}(x,y) dx dy$$

Example $g(x, y) = 3x, f_{x,y} = 4xy, 0 < x < 1, 0 < y < 1$ $E(g(x, y)) = \int_0^1 \int_0^1 3x \times 4xy \, dx \, dy$

Covariance of two random variables X and Y

$$Cov[x, y] = E[(x - E[x])(y - (E[y]))]$$
$$= E(XY) - E(X)E(Y)$$

If X and Y are independent, then:

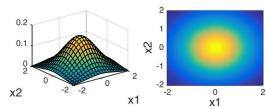
$$E(XY) = E(X)E(Y) \rightarrow Cov[x, y] = 0$$
$$Var[X + Y] = [E(X + Y)]^2 - E((X + Y)^2)$$
$$Var[X + Y] = Var[X] + Var[Y] + 2Cov[X, Y]$$

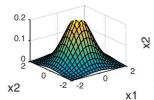
Multivariant Gaussian (Normal) distribution

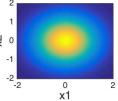
 $x \in \mathbb{R}^n$. Model $p(x_1), p(x_2), \dots etc.$ at the same time. Parameters : $\mu \in \mathbb{R}^n, \Sigma \in \mathbb{R}^{n \times n}$ (covariancematrix)

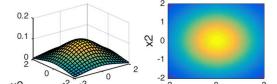
$$p(x;\mu,\Sigma) = \frac{1}{(2\pi)^{n/2}|\Sigma|^{\frac{1}{2}}} exp(-\frac{1}{2}(x-\mu)^{T}\Sigma^{-1}(x-\mu))$$

 $\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}^{\mathsf{T}}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ $\Sigma = \begin{array}{c} 0.7 \\ 0 \end{array}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ 0 0.7 $\Sigma = \begin{bmatrix} 1.5 \\ 0 \end{bmatrix}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ 0 1.5

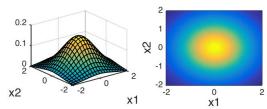


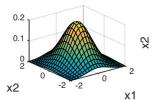


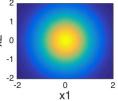


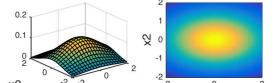


 $\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}^{\mathsf{T}}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ $\Sigma = \begin{array}{c} 0.6 \\ 0 \end{array}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ $\Sigma = \begin{bmatrix} 2 & 0 \\ 0 & 1 \end{bmatrix}^{\mathsf{T}}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$

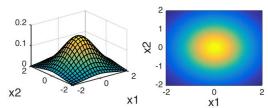


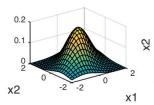


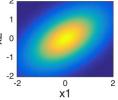


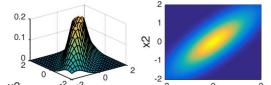


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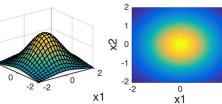


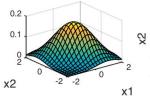


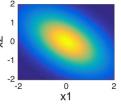


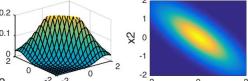


 $\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ 0.2 0.1 02 x2 $\Sigma = \begin{array}{c} 1\\ -0.5 \end{array}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ -0.5 0.2 0.1 02 -2 x2 $\Sigma = \begin{array}{c} 1\\ -0.8 \end{array}$ $\mu = \begin{bmatrix} 0 & 0 \end{bmatrix}^{\mathsf{T}}$ -0.8 0.2 0.1 02

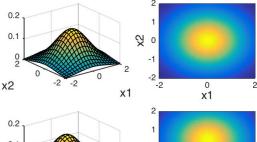


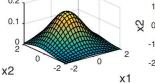


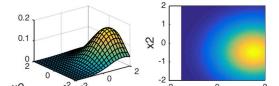




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-1 -2 -2

0

x1

2

Remember:

Let *B* be any event such that $P(B) \neq 0$. $P(A|B) := \frac{P(A \cap B)}{P(B)}$

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X,Y are RVs with the same probability space,

we have

$$P(X = x | Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$$

$$\mathbb{E}(X|Y=y) = \sum_{x} x \frac{P(X=x, Y=y)}{P(Y=y)}$$

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$\mathbb{E}[X|Y]$

$\mathbb{E}[X|Y]$

It is actually a random variable

$$\mathbb{E}[X|Y](y) = \mathbb{E}[X|Y=y]$$
 is a function of Y

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Law of Total Expectation

Let X, Y be RVs with the same probability space, then $\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X|Y]]$

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A brief proof of X,Y being discrete and finite

Law of Total Expectation

Let X, Y be RVs with the same probability space, then $\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X|Y]]$

A brief proof of X,Y being discrete and finite

$$\mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}[\sum_{x} xP(X = x|Y)]$$

= $\sum_{y} (\sum_{x} xP(X = x|Y = y))P(Y = y)$
= $\sum_{y} \sum_{x} xP(X = x, Y = y)$
= $\sum_{x} x(\sum_{y} P(X = x, Y = y))$
= $\sum_{x} xP(X = x)$
= $\mathbb{E}[X]$

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More Conditioned Bayes Rule

$$P(a|b,c) = rac{P(b|a,c)P(a|c)}{P(b|c)}$$

It is actually the same as the Bayes Rule:

$$P(a|b) = rac{P(b|a)P(a)}{P(b)}$$

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with a random variable c that all the probabilities are conditioned on.

More Conditioned Bayes Rule

A proof:

$$\frac{P(b|a,c)P(a|c)}{P(b|c)} = \frac{P(b,a,c)P(a|c)}{P(b|c)P(a,c)}$$
$$= \frac{P(b,a,c)P(a,c)}{P(b|c)P(a,c)P(c)}$$
$$= \frac{P(b,a,c)}{P(b|c)P(c)}$$
$$= \frac{P(b,a,c)}{P(b,c)}$$
$$= P(a|b,c)$$