

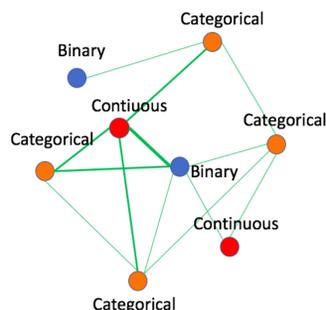


# Learning the Network Structure of Heterogeneous Data via Pairwise Exponential MRF

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## Motivation

- Typical Multivariate distributions (e.g. GMRF and Ising model) can model data on the homogeneous domain.
- But, the entities of real data can have heterogeneous domain (e.g. binary, category, and continuous domain).
- Need to define a proper multivariate distribution for heterogeneous data.
- Propose a subset of multivariate exponential family distribution, pairwise exponential Markov Random Fields, able to reveal the Markov (Network) structure across the entities.



## Model: Pairwise Exponential Markov Random Field

- For a Markov Random Fields described by  $G = (V, E)$  with  $|V| = p$ , a random vector  $X = \{X_1, \dots, X_p\}$  is defined over (heterogeneous) domains  $\mathcal{X} = \otimes_{r=1}^p \mathcal{X}_r$ .
- Let each node  $X_r$  have  $B_r(X_r)$  be ( $m_r$ -dimensional) node-potential and scalar base measure  $C_r(X_r)$ , by which we can design the conditional distribution on  $X_r$  given others  $X_{\setminus r}$ .

**Definition** A random vector  $X$  is defined as a PE-MRF if, for  $x = \{x_1, \dots, x_p\} \in \mathcal{X}$ , it follows the joint distribution

$$p(x; \theta) = \exp\left\{ \sum_{r=1}^p \theta_r^T B_r(x_r) + \sum_{s,t=1}^p \langle \Theta_{st}, B_t(x_t) B_s(x_s)^T \rangle_F + \sum_{r=1}^p C_r(x_r) - A(\theta) \right\}.$$

- $\theta$  is the natural parameter with node-parameter  $\theta_r \in \mathbb{R}^{m_r}$  and edge-parameter  $\Theta_{st} \in \mathbb{R}^{m_s \times m_t}$ .
- The log-partition function  $A(\theta)$  should be finite.

### Property.

- PE-MRF is exponential family where sufficient statistics and natural parameter interact linearly and nodes have a pairwise relationship.
- Include well-known distributions such as GMRF, Ising model, discrete model.
- The  $\{\Theta_{st}\}_{s,t=1}^p$  explicitly reveal underlying Markov structure.

## Estimation: Approximated Maximum Likelihood

- Maximum likelihood would be a typical approach but it entails intractable  $A(\theta)$ .
- Instead, replace  $A(\theta)$  with a tractable upperbound  $U(\theta)$

$$\underset{\theta}{\text{minimize}} \quad \underbrace{-\langle \theta, \overline{\mathbf{B}(\mathbf{x})} \rangle + U(\theta)}_{\text{approximated likelihood function}} + \underbrace{R_\lambda(\theta)}_{\text{group lasso}}.$$

- Here,  $\overline{\mathbf{B}(\mathbf{x})}$  is the averaged sufficient statistic  $B(x^{(i)})$  over the  $n$  samples. And the group lasso  $R_\lambda(\theta) = \lambda \sum_{s \neq t} w_{st} \|\Theta_{st}\|_F$  [2] encourages the  $st$ -th block, for every  $s \neq t$ , to be a zero matrix.
- Derive a convex upperbound  $U(\theta)$ . (Sketch: Express  $A(\theta)$  as an entropy  $H(X)$ , simplify it into  $\equiv H(B_1(X_1), \dots, B_p(X_p))$ , and use the fact that Gaussian distribution gives maximum entropy among any pdf under the same covariance matrix.)

**Theorem.** For a PE-MRF, the approximated negative maximum log-likelihood problem becomes

$$\min_{\Theta \in \mathcal{S}^{d+1}} \left\{ \langle \Theta, \overline{\mathbf{B}_{aug}(\mathbf{x})} \rangle_F - \log \det \Theta + R_\lambda(\Theta) \right\}$$

where  $d = \sum_{r=1}^p m_r$ ,  $\Theta$  is the augmented natural parameter of  $\theta$ ,  $\overline{\mathbf{B}_{aug}(\mathbf{x})}$  is augmented and shifted version of  $\overline{\mathbf{B}(\mathbf{x})}$ , and  $R_\lambda(\Theta) \equiv R_\lambda(\theta)$ .

- Call it *group graphical lasso* that extends the classic *graphical lasso* problem to general setting.

## Algorithm

- Formulate the problem below and solve via alternating direction method of multipliers (ADMM) [1]

$$\min_{\Theta, \mathbf{Z}} \langle \Theta, A \rangle_F - \log \det \Theta + \lambda_n \sum_{i \neq j} w_{ij} \|Z_{ij}\|_F.$$

- ADMM solves the augmented Lagrangian in iterative manner with respect to  $\Theta$ ,  $\mathbf{Z}$ , and  $\mathbf{U}$  where  $\mathbf{U}$  is the scaled dual variables.
- At  $k$ th iteration, all of updates have the following the closed form:  
 **$\Theta$ -Update.**  $\Theta^{k+1} := 1/2\eta Q(\Lambda + \sqrt{\Lambda^2 + 4\eta I})Q^T$  where  $\eta = \rho/n$  and  $Q\Lambda Q^T$  is the eigendecomposition of  $\eta(\mathbf{Z}^k - \mathbf{U}^k) - \overline{\mathbf{B}_{aug}(\mathbf{x})}$  [5].

**$\mathbf{Z}$ -Update.** For  $i, j \in \{1, \dots, p\}$ , compute  $Z_{ij, i \neq j} = \left(1 - \frac{\lambda_n w_{ij}}{\rho \|\Theta_{ij}^{k+1} + U_{ij}^k\|_F}\right) (\Theta_{ij}^{k+1} + U_{ij}^k)$  if  $\|\Theta_{ij}^{k+1} + U_{ij}^k\|_F \geq \lambda_n w_{ij} / \rho$  or 0 otherwise. And  $\mathbf{Z}^{k+1} := \Theta^{k+1} + \mathbf{U}^k$  for the rest elements.

**$\mathbf{U}$ -Update.**  $\mathbf{U}^{k+1} := \mathbf{U}^k + \Theta^{k+1} - \mathbf{Z}^{k+1}$ .

## Edge Recovery Consistency

- Assume a PE-MRF  $X$  satisfies the graph condition (Incoherence condition on the Hessian matrix [3] and graph structure) and boundness condition on sufficient statistics.

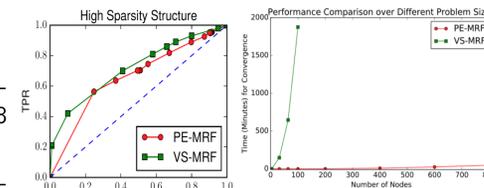
**Theorem.** For  $\lambda_n \geq \kappa_1 \sqrt{\frac{\log(m_{max} p)}{n}}$ , let  $\hat{\Theta}$  be the (unique) solution of *group graphical lasso*. If the number of samples  $n \geq \kappa_2 \log m_{max} p$ , then the estimated edge  $E(\hat{\Theta}) = \{(s, t) \mid \|\hat{\Theta}_{ij}\|_2 \geq \kappa_3 \lambda_n\}$  can exactly recover the real edge set  $E$  with probability at least  $1 - e^{-cn}$ .

- Here  $\kappa_1, \kappa_2$  and  $\kappa_3$  depend on the  $\{m_r\}, \{w_{st}\}$  and other parameters defined in assumptions, and  $c$  is some universal constant.

## Experiments

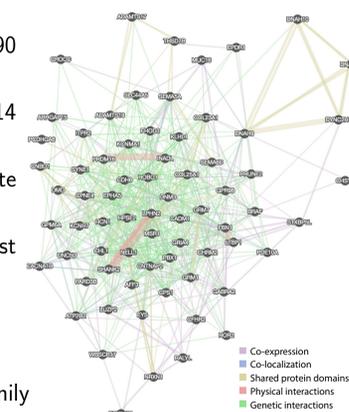
### Synthetic Data.

- Experiment on (synthetic) heterogeneous network with 32 nodes: 8 Bernoulli, 8 gamma, 8 Gaussian, and 8 three-dimensional Dirichlet.
- Plot the ROC curves for edge recovery percentage, compared with VS-MRF [4].
- For inferring the 100-node Markov network, our PE-MRF solver can takes **under 30 seconds** (whereas VS-MRF takes **over 31 hours**).



### Heterogeneous Genomic Networks.

- Level III public data from The Cancer Genome Atlas for 290 breast cancer patients.
- Contains expression profiles for 500 genes (Gaussian) and 314 miRNAs (Poisson).
- Consider a 65-core of the gene-gene subnetwork to demonstrate the its utility.
- For validation, observe how many gold standard edges exist between 65-core genes from external data.



## Conclusion

- Propose a PE-MRF model, a subclass of the exponential family that is well-suited for heterogeneous multivariate distributions.
- Formulate an approximated maximum likelihood problem by deriving upper bound on the log partition function.
- Develop an  $O(p^3)$  ADMM algorithm with closed-form updates.
- The estimator guarantees to recovery underlying Markov structure consistently.
- Our results, as well as the widespread applications with heterogeneous data sources, lead to many extensions of this work. For example, instead of inferring a single PE-MRF network, we could use the time-stamped observations to estimate a time-varying network because it is possible that Markov structure changes over time.

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