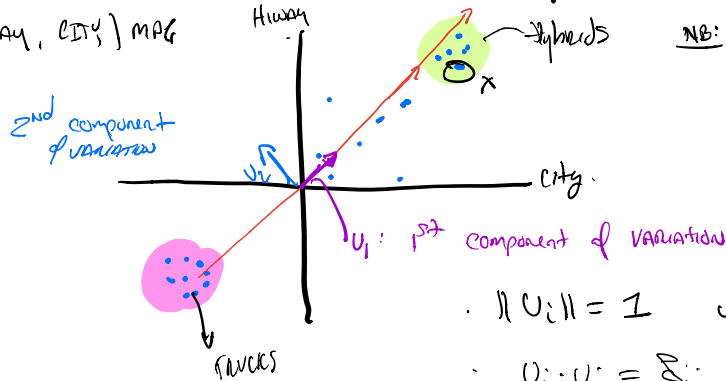


PCA & ICA

PCA: Principal Component Analysis

GIVEN (Highway, City) MAP



NB: WE CENTERED DATA

$$x^{(i)} \mapsto x^{(i)} - \mu$$

$$\mu = \frac{1}{n} \sum_{i=1}^n x^{(i)}$$

$$\cdot \|u_i\| = 1 \quad \text{unit length}$$

$$\cdot u_i \cdot u_j = \delta_{ij} \quad (\text{orthonormal})$$

• u_1 - "How good is the mpg"

• u_2 - "variation in city/highway from 'good'"

$$x = \alpha_1 u_1 + \alpha_2 u_2$$

$$x^{(i)} = \alpha_1^{(i)} u_1 + \alpha_2^{(i)} u_2$$

today How we find directions

Think about dimension $k=2 \rightarrow 10$

Preprocessed

GIVEN $x^{(1)} \dots x^{(n)} \in \mathbb{R}^d$

1. CENTER THE DATA $x^{(i)} \mapsto x^{(i)} - \mu \quad \mu = \frac{1}{n} \sum x^{(i)}$

2. MAY NEED TO RESCALE COMPONENTS -- "FEET PER GALLON" "MPG"

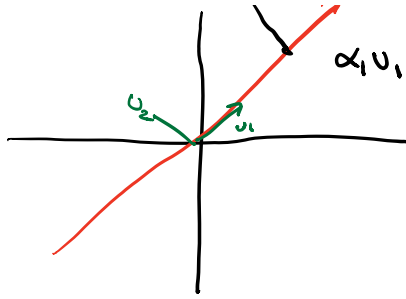
"WHITEN" compute sample variance ...

$$x_j^{(i)} \mapsto \frac{(x_j^{(i)} - \mu_j)}{\sigma_j}$$

WE WILL ASSUME DATA IS PREPROCESSED.

PCA AS OPTIMIZATION

* $u_i \in \{t u_i : t \in \mathbb{R}\}$ line corresponds to u_i



How do you find closest point to the line generated by u_1 ?

$$\alpha_1 = \operatorname{argmin}_{\alpha} \|x - \alpha u_1\|^2$$

$$= \operatorname{argmin}_{\alpha} \underbrace{\|x\|^2 + \alpha^2 \|u_1\|^2 - 2\alpha (u_1 \cdot x)}_{g(\alpha)}$$

$$\nabla g(\alpha) = 2\alpha - 2(u_1 \cdot x) = 0 \Rightarrow \alpha = u_1 \cdot x$$

Generalize: $u_1, \dots, u_k \in \mathbb{R}^d$ AND $x \in \mathbb{R}^d$ ($u_i \cdot u_j = \delta_{i,j}$)

$$\operatorname{Argmin}_{\alpha_1, \dots, \alpha_k} \|x - \sum_{j=1}^k \alpha_j u_j\|^2 \dots$$

Hence $\alpha_j = u_j \cdot x$

$$\|x - \sum_{j=1}^k \alpha_j u_j\|^2 \leftarrow \text{RESIDUAL}$$

We can find PCA by either:

1. MAXIMIZE Projected subspace
2. MINIMIZE Residuals

$$\max_{\substack{U \in \mathbb{R}^d \\ \|U\| = 1}} \frac{1}{n} \sum_{i=1}^n (x^{(i)} \cdot U)^2$$

Solve this optimization problem we need some facts.

LET A be symmetric & square then

$$A = U \Lambda U^T \text{ in which}$$

$$\cdot U U^T = U^T U = I \text{ (ORTHONORMAL BASIS)}$$

$\cdot \Lambda$ IS DIAGONAL MATRIX

$$\Lambda_{ii} = \lambda_i \text{ AND } \lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d \text{ by convention}$$

\rightarrow eigenvalues

RECALL If $x = \sum_{i=1}^d \alpha_i u_i$ where $[u_1, \dots, u_d] = U$

$$\begin{aligned} Ax &= U \Lambda U^T x = U \Lambda \sum_{i=1}^d \alpha_i e_i & (u_i \cdot u_j = \delta_{ij}) \\ &= U \sum_{i=1}^d \alpha_i \lambda_i e_i & \text{if } i=j \Rightarrow \delta_{ij}=1 \\ &= \sum_{i=1}^d \alpha_i \lambda_i u_i & \text{else } 0 \end{aligned}$$

fix i , and let $c \in \mathbb{R} \neq 0$

$$x = c u_i \text{ eigenvectors } Ax = \lambda_i x$$

$$\text{MAX}_{U \in \mathbb{R}^d} \|U\|=1$$

$$\frac{1}{n} \sum_{i=1}^n (x^{(i)} \cdot u)^2$$

$$\Leftrightarrow \text{MAX}_{U \in \mathbb{R}^d} U^T A U$$

$\|U\|=1$

$$A = \frac{1}{n} \sum_{i=1}^n x^{(i)} x^{(i)T}$$

$$\Leftrightarrow \text{MAX}_{\alpha \in \mathbb{R}^d} \sum_{i=1}^d \alpha_i^2 \lambda_i$$

$\|x\|^2=1$

WHAT SHOULD WE PICK TO MAXIMIZE? $\alpha_1=1 \dots \alpha_2=\dots=\alpha_d=0$

IS IT UNIQUE? $\lambda_1 = \lambda_2$ WHAT HAPPENS? (PCA "loopy instability")

$$\lambda_1 > \lambda_2 \Rightarrow \text{UNIQUE}$$

u_1 IS THE PRINCIPAL EIGENVECTOR

WHAT IF WE WANT THE TOP-K SVD VECTORS?

$$u_1, \dots, u_k \text{ BECAUSE } \lambda_1 \geq \dots \geq \lambda_k$$

$$x^{(i)} \mapsto \sum_{j=1}^k (x^{(i)} \cdot u_j) u_j$$

$j=$ keep these knowers.

How do we pick k ?

ONE Approach "Amount of Explained VARIANCE"

$$\frac{\sum_{i=1}^k \lambda_i}{\sum_{j=1}^n \lambda_j} \geq 0.9 \quad \lambda_i \geq 0$$

Lucky Instability $\lambda_k = \lambda_{k+1} = \lambda_{k+2}$ Are you were lucky for k ?

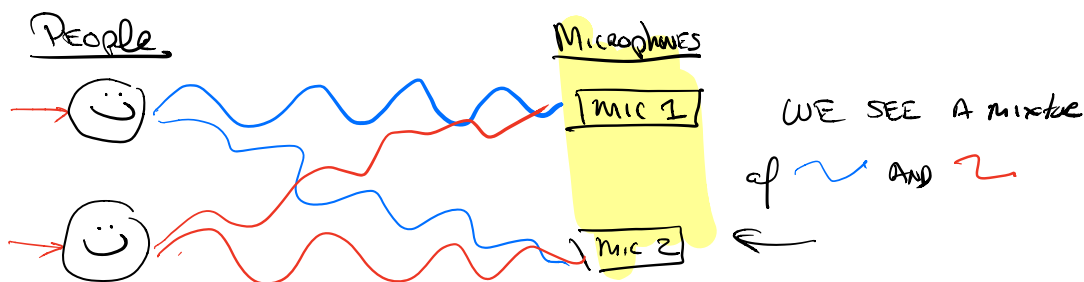
Pick any 2 of these! \Rightarrow different bases

PCA · Dimensionality reduction technique

- MAIN IDEA IS Project on Subspace
- Nice story! Contrast w/ EM

ICA Independent Component Analysis

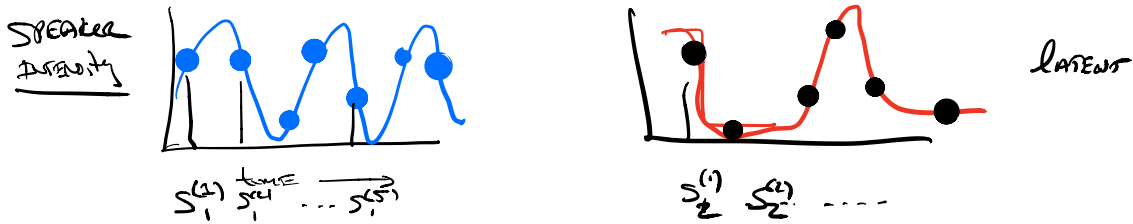
- High-level Story
- key facts AND likelihood
- model



SPEAKER SEPARATION.

SPEAKER $S_1^{(t)}, S_2^{(t)}$

DATA $X_1^{(t)}, X_2^{(t)}$



$S_j^{(t)}$ IS THE INTENSITY AT TIME t OF SPEAKER j

WE DO NOT OBSERVE $S_j^{(t)}$ ONLY $x_j^{(t)}$ - THE MICROPHONE

MODEL $x_j^{(t)} = a_{j1} S_1^{(t)} + a_{j2} S_2^{(t)} \leftarrow$ HIDDEN

"MICROPHONE j SEES A MIXTURE OF THE SPEAKER INTENSITIES"

OBSERVED $\leftarrow x^{(t)} = A s^{(t)} \rightarrow$ LATENT

FOR SIMPLICITY, $d \geq$ THE NUMBER OF SPEAKERS & MICROPHONES

GIVEN: $x^{(1)}, \dots, x^{(n)} \in \mathbb{R}^d$

DO: FIND $s^{(1)}, \dots, s^{(n)} \in \mathbb{R}^d$

AND $A \in \mathbb{R}^{d \times d}$ S.T. $x^{(t)} = A s^{(t)}$

WE CALL A THE MIXING MATRIX AND $W = A^{-1}$ UNMIXING MATRIX

WRITE $N = \begin{bmatrix} w_1^T \\ \vdots \\ w_d^T \end{bmatrix}$ SO THAT $S_j^{(t)} = w_j \cdot x^{(t)}$

WE CENTER THE DATA $x^{(t)} \mapsto x^{(t)} - \mu$ $\mu = \frac{1}{n} \sum x^{(t)}$

• A DOES NOT VARY w/ TIME, A IS FULL RANK

• THERE ARE SOME INHERENT AMBIGUITY:

• WE CAN'T DETERMINE SPEAKER ID

• CAN'T DETERMINE INTENSITY (ABSOLUTE)

$$x^{(i)} = A s^{(i)}$$

$$= (cA) (c^{-1} s^{(i)}) \text{ for any } c \neq 0.$$

Surprising SPEAKERS CANNOT BE GAUSSIAN

$$x^{(i)} = A s^{(i)} \quad s^{(i)} \sim N(0, I)$$

$$\Rightarrow x^{(i)} \sim N(0, AA^T) \quad UU^T = I$$

$$AUU^T A^T = AA^T$$

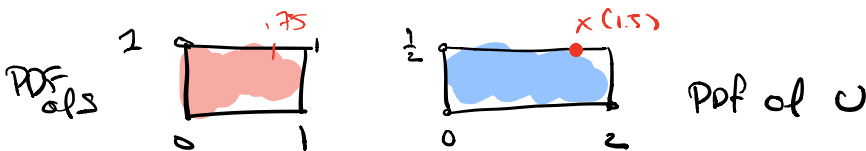
Nevertheless, we can recover something meaningful

Algorithm: Just MLE w/ Gradient Descent.

Detail: Density under linear transform

Ex: SN Uniform $[0, 1]$ $U = 2S$ WHAT'S THE PDF of U ?

$$\text{transformed } P_U(x) = P_S\left(\frac{x}{2}\right) \text{ (see left)}$$



$$P_S(x) \begin{cases} 1 & \text{if } x \in [0, 1] \\ 0 & \text{o.w.} \end{cases} \quad P_U(x) = \underline{P_S\left(\frac{x}{2}\right)} \cdot \frac{1}{2}$$

THE KEY ISSUE IS THE NORMALIZATION CONSTANT

A square & invertible, $U = AS$ $S \sim \text{PDF of } P_S$

$$P_U(x) = P_S(A^{-1}x) |\det(A^{-1})|$$

$$= \underline{P_S(Wx)} |\det(W)|$$

CHANGE of VARIABLES formula



volume $\text{vol}(B)$

$$\text{Vol}(AB) = |\det(A)| \text{vol}(B)$$

FROM HERE ICA IS MLE!

$$\text{Latent} \rightarrow P(s) = \prod_{j=1}^J P_j(s_j)$$

"SOURCES INDEPENDENT"

AND HAVE DISTRIBUTION"

$$\text{OBSERVED} \rightarrow P(x) = \prod_{j=1}^J P_j(w_j \cdot x) |\det(w)|$$

Key technical trick NOT ROTATIONALLY SYMMETRIC

SET $B(x) \propto g'(x)$ for $g(x) = (1 + e^{-x})^{-1}$

$$l(w) = \sum_{t=1}^n \sum_{j=1}^J \log g'(w_j \cdot x^{(t)}) + \log |\det(w)|$$

- log det w
- USE GD & you're done

RECAP:

- SAME PIA. WORKHOUSE DIMENSIONALITY REDUCTION
- ICA - KEY IDEAS. ITANORLY SYMMETRY.